1. Faults occur in a roll of material at a rate of $\lambda$ per $\mathrm{m}^{2}$. To estimate $\lambda$, three pieces of material of sizes $3 \mathrm{~m}^{2}, 7 \mathrm{~m}^{2}$ and $10 \mathrm{~m}^{2}$ are selected and the number of faults $X_{1}, X_{2}$ and $X_{3}$ respectively are recorded.

The estimator $\hat{\lambda}$, where

$$
\hat{\lambda}=k\left(X_{1}+X_{2}+X_{3}\right)
$$

is an unbiased estimator of $\lambda$.
(a) Write down the distributions of $X_{1}, X_{2}$ and $X_{3}$ and find the value of $k$.
(b) Find $\operatorname{Var}(\hat{\lambda})$.

A random sample of $n$ pieces of this material, each of size $4 \mathrm{~m}^{2}$, was taken. The number of faults on each piece, $Y$, was recorded.
(c) Show that $\frac{1}{4} \bar{Y}$ is an unbiased estimator of $\lambda$.
(d) Find $\operatorname{Var}\left(\frac{1}{4} \bar{Y}\right)$.
(3)
(e) Find the minimum value of $n$ for which $\frac{1}{4} \bar{Y}$ becomes a better estimator of $\lambda$ than $\hat{\lambda}$.
2. The weights of the contents of jars of jam are normally distributed with a stated mean of 100 g . A random sample of 7 jars was taken and the contents of each jar, $x$ grams, was weighed. The results are summarised by the following statistics.

$$
\sum x=710.9, \quad \sum x^{2}=72219.45 .
$$

Test at the $5 \%$ level of significance whether or not there is evidence that the mean weight of the contents of the jars is greater than 100 g . State your hypotheses clearly.
(Total 8 marks)
3. An engineer decided to investigate whether or not the strength of rope was affected by water. A random sample of 9 pieces of rope was taken and each piece was cut in half. One half of each piece was soaked in water overnight, and then each piece of rope was tested to find its strength. The results, in coded units, are given in the table below

| Rope no. | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Dry rope | 9.7 | 8.5 | 6.3 | 8.3 | 7.2 | 5.4 | 6.8 | 8.1 | 5.9 |
| Wet rope | 9.1 | 9.5 | 8.2 | 9.7 | 8.5 | 4.9 | 8.4 | 8.7 | 7.7 |

Assuming that the strength of rope follows a normal distribution, test whether or not there is any difference between the mean strengths of dry and wet rope. State your hypotheses clearly and use a $1 \%$ level of significance.
(Total 8 marks)
4. A certain vaccine is known to be only $70 \%$ effective against a particular virus; thus $30 \%$ of those vaccinated will actually catch the virus. In order to test whether or not a new and more expensive vaccine provides better protection against the same virus, a random sample of 30 people were chosen and given the new vaccine. If fewer than 6 people contracted the virus the new vaccine would be considered more effective than the current one.
(a) Write down suitable hypotheses for this test.
(b) Find the probability of making a Type I error.
(c) Find the power of this test if the new vaccine is
(i) $80 \%$ effective,
(ii) $90 \%$ effective.

An independent research organisation decided to test the new vaccine on a random sample of 50 people to see if it could be considered more than $70 \%$ effective. They required the probability of a Type I error to be as close as possible to 0.05 .
(d) Find the critical region for this test.
(e) State the size of this critical region.
(f) Find the power of this test if the new vaccine is
(i) $80 \%$ effective,
(ii) $90 \%$ effective.
(g) Give one advantage and one disadvantage of the second test.
5. Gill, a member of the accounts department in a large company, is studying the expenses claims of company employees. She assumes that the claims, in $£$, follow a normal distribution with mean $\mu$ and variance $\sigma^{2}$. As a first stage in her investigation she took the following random sample of 10 claims.
30.85, 99.75, 142.73, 223.16, 75.43, 28.57, 53.90, 81.43, 68.62, 43.45.
(a) Find a $95 \%$ confidence interval for $\mu$.

The chief accountant would like a 95\% confidence interval where the difference between the upper confidence limit and the lower confidence limit is less than 20.
(b) Show that $\frac{\sigma^{2}}{n}<26.03$ (to 2 decimal places), where $n$ is the size of the sample required to achieve this.

Gill decides to use her original sample of 10 to obtain a value for $\sigma^{2}$ so that the chance of her value being an underestimate is 0.01 .
(c) Find such a value for $\sigma^{2}$.
(d) Use this value for $\sigma^{2}$ to estimate the size of sample the chief accountant requires.
6. An educational researcher is testing the effectiveness of a new method of teaching a topic in mathematics. A random sample of 10 children were taught by the new method and a second random sample of 9 children, of similar age and ability, were taught by the conventional method. At the end of the teaching, the same test was given to both groups of children.

The marks obtained by the two groups are summarised in the table below.

|  | New method | Conventional method |
| :--- | :---: | :---: |
| Mean $(\bar{X})$ | 82.3 | 78.2 |
| Standard deviation $(s)$ | 3.5 | 5.7 |
| Number of students $(n)$ | 10 | 9 |

(a) Stating your hypotheses clearly and using a $5 \%$ level of significance, investigate whether or not
(i) the variance of the marks of children taught by the conventional method is greater than that of children taught by the new method,
(ii) the mean score of children taught by the conventional method is lower than the mean score of those taught by the new method.
[In each case you should give full details of the calculation of the test statistics.]
(b) State any assumptions you made in order to carry out these tests.
(c) Find a 95\% confidence interval for the common variance of the marks of the two groups.
7. A statistics student is trying to estimate the probability, $p$, of rolling a 6 with a particular die. The die is rolled 10 times and the random variable $X_{1}$ represents the number of sixes obtained. The random variable $R_{1}=\frac{X_{1}}{10}$ is proposed as an estimator of $p$.
(a) Show that $R_{1}$ is an unbiased estimator of $p$.

The student decided to roll the die again $n$ times $(n>10)$ and the random variable $X_{2}$ represents the number of sixes in these $n$ rolls. The random variable $R_{2}=\frac{X_{2}}{n}$ and the random variable $Y=$ $\frac{1}{2}\left(R_{1}+R_{2}\right)$.
(b) Show that both $R_{2}$ and $Y$ are unbiased estimators of $p$.
(c) Find Var $\left(R_{2}\right)$ and $\operatorname{Var}(Y)$.
(d) State giving a reason which of the 3 estimators $R_{1}, R_{2}$ and $Y$ are consistent estimators of $p$.
(e) For the case $n=20$ state, giving a reason, which of the 3 estimators $R_{1}, R_{2}$ and $Y$ you would recommend.

The student's teacher pointed out that a better estimator could be found based on the random variable $X_{1}+X_{2}$.
(f) Find a suitable estimator and explain why it is better than $R_{1}, R_{2}$ and $Y$.

1. (a) $X_{1} \sim \operatorname{Po}(3 \lambda)$

$$
\begin{array}{rlrl}
X_{2} & \sim \operatorname{Po}(7 \lambda) & \text { M1 } \\
X_{3} & \sim \operatorname{Po}(10 \lambda) & \\
\mathrm{E}(\hat{\lambda}) & =k\left[\mathrm{E}\left(X_{1}\right)+\mathrm{E}\left(X_{2}\right)+\mathrm{E}\left(X_{3}\right)\right] & \text { M1 } \\
& =20 \lambda k & & \text { M1 } \\
\hat{\lambda} \text { unbiased therefore } 20 \lambda k=\lambda & & \text { A1 }
\end{array}
$$

## Note

M1 all 3 needed. Poisson and mean
M1 adding their means
M1 putting their $\mathrm{E}(\hat{\lambda})=\lambda$
A1 cao
(b) $\operatorname{Var}(\hat{\lambda})=\frac{1}{20^{2}} \operatorname{Var}\left(X_{1}+X_{2}+X_{3}\right) \quad$ M1

$$
\begin{array}{ll}
=\frac{1}{20^{2}}(3 \lambda+7 \lambda+10 \lambda) & \text { M1 } \\
=\frac{\lambda}{20} & \text { A1ft }
\end{array}
$$

## Note

M1 use of $k^{2} \operatorname{Var}\left(X_{1}+X_{2}+X_{3}\right)$
M1 using their means from part(a) as Variances and adding together A1 cao
(c) $\quad Y \sim \operatorname{Po}(4 \lambda)$

$$
\mathrm{E}\left(\frac{1}{4} \bar{Y}\right)=\frac{1}{4} \times 4 \lambda=\lambda \text { therefore unbiased }
$$

## Note

M1 use of $4 \lambda$
A1 cso plus conclusion. Accept working out bias to $=0$
(d) $\operatorname{Var}\left(\frac{1}{4} \bar{Y}\right)=\frac{1}{16} \times \frac{4 \lambda}{n}$

$$
=\frac{\lambda}{4 n}
$$

## Note

M1 $\frac{1}{16} \times \operatorname{Var} \bar{Y}$
B1 for $\operatorname{Var} \bar{Y}=\frac{4 \lambda}{n}$
A1 cao
(e) $\frac{\lambda}{4 n}<\frac{\lambda}{20} \quad$ M1

$$
n>5 \text { therefore } n=6
$$

A1 2

## Note

M1 for $\operatorname{Var}\left(\frac{1}{4} \bar{Y}\right)>\operatorname{Var}(\hat{\lambda})$
A1 n = 6
2. $\mathrm{H}_{0}: \mu=100, \mathrm{H}_{1}: \mu>100$

B1

$$
\begin{array}{ll}
\bar{x} & =\frac{710.9}{7}=101.5571 \ldots ; s^{2}=\frac{72219.45-\frac{(710.9)^{2}}{7}}{6} \\
s^{2} & =3.746 \ldots \text { or } s=1.9355
\end{array} \quad \text { B1, M1 }
$$

test statistic $t=\frac{101.557-100}{\frac{1.936}{\sqrt{7}}}=$ awrt 2.13
$t_{6} 5 \%$ 1-tail critical value $=1.943$
Significant result. Reject $\mathrm{H}_{0}$, there is evidence that the mean weight is more than 100 g .

M1 A1

B1 ft
3. $D=$ dry—wet $\mathrm{H}_{0}: \mu_{D}=0, \mathrm{H}_{1}: \mu_{D} \neq 0$

B1
$d: 0.6,-1,-1.9,-1.4,-1.3,0.5,-1.6,-0.6,-1.8$ M1
$\bar{d}:-\frac{8.5}{9}=-0.9 \dot{4}, s_{d}^{2}=\frac{15.03-9 \times(\bar{d})^{2}}{8}=0.87527 \ldots$
$t=\frac{-0.9 \dot{4}}{\frac{s_{d}}{\sqrt{9}}}=$ awrt -3.03 M1, A1 B1

Not significant - insufficient evidence of a difference between mean strength

A1 ft 8
4.
(a) $\mathrm{H}_{0}: p=0.3($ or 0.7$) \quad \mathrm{H}_{1}: p<0.3($ or $>0.7)$
(b) Let $X=$ number who contract virus. Under $\mathrm{H}_{0} \quad X \sim \mathrm{~B}(30,0.3)$ $\mathrm{P}($ Type I error $)=\mathrm{P}(X<6 \mid p=0.30)=\mathrm{P}(X \leq 5)=0.0766 \quad$ M1 A1 2
(c) (i) Power $=\mathrm{P}(Y \leq 5 \mid Y \sim \mathrm{~B}(30,0.2))=0.4275$
(ii) $\quad$ Power $=\mathrm{P}(Y \leq 5 \mid Y \sim \mathrm{~B}(30,0.1))=0.9268$

A1 3
(d) Let $C=$ number who contract virus. Under $\mathrm{H}_{0} \quad C \sim \mathrm{~B}(50,0.3)$

We require $c$ such that $\mathrm{P}(C \leq c) \approx 0.05 \quad$ M1
$\mathrm{P}(C \leq 10)=0.0789, \quad \mathrm{P}(C \leq 9)=0.0402 \quad \therefore$ critical region is $C \leq 9 \quad \mathrm{~A} 1 \quad 2$
(e) Size $=0.0402 \quad$ B1
(f) (i) Power $=\mathrm{P}(D \leq 9 \mid D \sim \mathrm{~B}(50,0.2))=0.4437 \quad$ B1
(ii) Power $=\mathrm{P}(D \leq 9 \mid D \sim \mathrm{~B}(50,0.1))=0.9755 \quad$ B1 2
(g) Advantage: second test is more powerful

Disadvantage: second test involves greater sample size, B1
$\therefore$ more expensive or takes longer $\quad$ B1 2
5. (a) $\bar{x}=\frac{847.89}{10}=84.79 ; \quad s_{x}^{2}=\frac{103712.6151-(847.89)^{2} / 10}{9} \quad$ B1
$s_{x}^{2}=3535.6522 \ldots$ B1
or $s_{X}=59.461 \ldots$
2.262

B1
$95 \%$ confidence interval for $\mu=84.79 \pm 2.262 \times \frac{59.461}{\sqrt{10}}=(42.25,127.33)$
accept $(42.3,127.3)$
M1, A1, A16
(b) $95 \%$ confidence interval $\bar{x} \pm 1.96 \frac{\sigma}{\sqrt{n}}$

$$
\text { so chief accountant requires } 1.96 \frac{\sigma}{\sqrt{n}}<10 \quad \text { M1 A1 }
$$

i.e. $\frac{\sigma^{2}}{n}<\left(\frac{10}{1.96}\right)^{2}=26.0308 \ldots=26.03$ (2 d.p.)

A1 cso 3
(c) Require the upper confidence limit of $98 \%$ confidence interval for $\sigma^{2}$

$$
\begin{aligned}
& \chi_{9}^{2}=2.088 \text {; i.e. } \frac{9 s^{2}}{\sigma^{2}}>2.088, \Rightarrow \sigma^{2}<15239.88 \ldots \\
& \text { awrt } 15240 \\
& \text { B1; M1, A13 } \\
& \text { (d) Substitute into part (b), } n>\frac{15240}{26.03} \Rightarrow n=586
\end{aligned}
$$

6. (a) (i) $\mathrm{H}_{0}: \sigma_{C}^{2}=\sigma_{N}^{2}, \mathrm{H}_{1}: \sigma_{C}^{2}>\sigma_{N}^{2}$ B1
$\frac{s_{C}^{2}}{s_{N}^{2}}=\frac{5.7^{2}}{3.5^{2}}=2.652 \ldots ; \quad \quad \mathrm{F}_{8,9}(5 \%)$ critical value $=3.23 \mathrm{M} 1 ; \mathrm{B} 1$
Not significant so do not reject $\mathrm{H}_{0}$ - insufficient evidence
A1 ft 4 that variance using conventional method is greater
(ii) $\mathrm{H}_{0}: \mu_{N}=\mu_{C}$
$\mathrm{H}_{1}: \mu_{N}>\mu_{C} \quad \mathrm{~B} 1$
$s^{2}=\frac{8 \times 5.7^{2}+9 \times 3.5^{2}}{17}=\frac{370.17}{17}=21.774 \ldots$
Test statistic $t=\frac{82.3-78.2}{\sqrt{21.774 \ldots\left(\frac{1}{9}+\frac{1}{10}\right)}}=1.9122 \ldots$
awrt 1.91
M1 A1
$t_{17}(5 \%)$ 1-tail critical value $=1.740$
Significant - reject $\mathrm{H}_{0}$. There is evidence that new style leads to an increase in mean
(b) Assumed population of marks obtained were normally distributed

B1 1
(c) Unbiased estimate of common variance is $s^{2}$ in (ii)

$$
\begin{array}{lc}
7.564<\frac{17 s^{2}}{\sigma^{2}}<30.191 & \text { B1 M1 B1 } \\
\sigma^{2}>\frac{17 \times 21.774 \ldots}{30.191}=12.3 \text { (1 d.p.) } & \text { A1 } \\
\sigma^{2}<\frac{17 \times 21.774 \ldots}{7.564}=48.9 \text { (1 d.p.) } & \text { A1 } 5
\end{array}
$$

7. (a) $\quad X_{1} \sim \mathrm{~B}(10, p) \quad \therefore \mathrm{E}\left(X_{1}\right)=10 p \Rightarrow \mathrm{E}\left(R_{1}\right)=\mathrm{E}\left(\frac{X_{1}}{10}\right)=\frac{10 p}{10}=p$

B1 1
(b) $\quad X_{2} \sim \mathrm{~B}(n, p) \quad \therefore \mathrm{E}\left(X_{2}\right)=n p \Rightarrow \mathrm{E}\left(R_{2}\right)=\mathrm{E}\left(\frac{X_{2}}{n}\right)=\frac{n p}{n}=p$ B1 $\mathrm{E}(Y)=\mathrm{E}\left(\frac{1}{2}\left[R_{1}+R_{2}\right]\right)=\frac{1}{2}\left[\mathrm{E}\left(R_{1}\right)+\mathrm{E}\left(R_{2}\right)\right]=\frac{1}{2}[p+p]=p$ B1 2
(c) $\operatorname{Var}\left(R_{2}\right)=\frac{1}{n^{2}} \operatorname{Var}\left(X_{2}\right)=\frac{n p(1-p)}{n^{2}}=\frac{p(1-p)}{n}$ B1

$$
\begin{align*}
& \operatorname{Var}\left(R_{1}\right)=\frac{p(1-p)}{10} \quad \therefore \operatorname{Var}(Y)=\frac{1}{4}\left[\operatorname{Var}\left(R_{1}\right)+\operatorname{Var}\left(R_{2}\right)\right], \\
& \quad=\frac{1}{4}\left[\frac{p(1-p)}{10}+\frac{p(1-p)}{n}\right]
\end{align*}
$$

(d) Since $\operatorname{Var}\left(R_{2}\right)=\frac{p(1-p)}{n} \rightarrow 0$ as $n \rightarrow \infty, \therefore R_{2}$ is consistent $\quad$ M1, A1 2
(e) $\operatorname{Var}\left(R_{1}\right)=\frac{p(1-p)}{10}>\frac{p(1-p)}{20}=\operatorname{Var}\left(R_{2}\right)$
$\operatorname{Var}(Y)=\frac{p(1-p)}{4}\left[\frac{1}{10}+\frac{1}{20}\right]=\frac{p(1-p)}{80} \times 3<\operatorname{Var}\left(R_{2}\right)$
Since all 3 are unbiased, we select the one with minimum
A1 2 variance, i.e. $Y$
(f) $\quad X_{1}+X_{2} \sim \mathrm{~B}(n+10, p)$ so consider $\frac{X_{1}+X_{2}}{n+10}$

$$
\mathrm{E}\left(\frac{X_{1}+X_{2}}{n+10}\right)=\frac{(n+10) p}{(n+10)}=p
$$

(show unbiased)
$\operatorname{Var}\left(\frac{X_{1}+X_{2}}{n+10}\right)=\frac{p(1-p)}{n+10}$
(find variance)
M1

$$
\frac{p(1-p)}{n+10}<\frac{p(1-p)}{10} \quad \therefore \text { always better than } R_{1}
$$

$\frac{p(1-p)}{n+10}<\frac{p(1-p)}{n} \quad \therefore$ always better than $R_{2}$
$\frac{p(1-p)}{n+10}<\frac{p(1-p)}{4}\left[\frac{n+10}{10 n}\right]$
$\Leftrightarrow \quad 40 n<100+20 n+n^{2}$
$\Leftrightarrow \quad 0<10^{2}-20 n+n^{2}$
$\Leftrightarrow \quad 0<(10-n)^{2}$
Show better than Y
Use of $\mathrm{n}=20$ acceptable M1
$\therefore \frac{X_{1}+X_{2}}{n+10}$ is unbiased and always has smaller variance A1 cso
6

1. This question proved to be the most challenging question for many candidates. Few candidates wrote down the distributions of $X_{1}, X_{2}$ and $X_{3}$ in part (a) and those who tried were unable to do so accurately. The Normal and Binomial distributions were commonly seen. This aside candidates were then able to progress and gain at least two marks in part (a).
In part (b) the main error was not to use their means from part (a). Even the candidates who correctly identified the Poison introduced a variety of variances including $\sigma^{2}$.

In part (c) and (d) the candidates who knew $\mathrm{E}(\bar{Y})=\mu$ and $\operatorname{Var}(\bar{Y})=\frac{\sigma^{2}}{n}$ gained full marks.
In part (e) the majority of candidates used $\operatorname{Var}\left(\frac{1}{4} \bar{Y}\right)<\operatorname{Var}(\hat{\lambda})$ although it was not always clear from their working that this was the case with many writing $\operatorname{Var}\left(\frac{1}{4} \bar{Y}\right)=\operatorname{Var}(\hat{\lambda})$ and simply solving the equation.
2. No Report available for this question.
3. No Report available for this question.
4. No Report available for this question.
5. No Report available for this question.
6. No Report available for this question.
7. No Report available for this question.

